

Feb 10

Def: X is a multivariate random vector.

i.e.

$$X = (x_1, \dots, x_n)^T$$

$$E(X) = \begin{bmatrix} E(x_1) \\ E(x_2) \\ \vdots \\ E(x_n) \end{bmatrix}$$

Ex: Let $(x_1, x_2, x_3)^T$ is a trivariate RV.

The covariance matrix of $(x_1, x_2, x_3)^T$ is

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defined as

$$\begin{matrix} & & x_2 & x_3 \\ & & & \\ x_1 & & & \\ x_2 & & & \\ x_3 & & & \end{matrix} \begin{pmatrix} \text{Cov}(x_1, x_1) & \text{Cov}(x_1, x_2) & \text{Cov}(x_1, x_3) \\ \text{Cov}(x_2, x_1) & \text{Cov}(x_2, x_2) & \text{Cov}(x_2, x_3) \\ \text{Cov}(x_3, x_1) & \text{Cov}(x_3, x_2) & \text{Cov}(x_3, x_3) \end{pmatrix}$$

$$= \begin{pmatrix} \text{Var}(x_1) & \text{Cov}(x_1, x_2) & \text{Cov}(x_1, x_3) \\ \text{Cov}(x_1, x_2) & \text{Var}(x_2) & \text{Cov}(x_2, x_3) \\ \text{Cov}(x_1, x_3) & \text{Cov}(x_2, x_3) & \text{Var}(x_3) \end{pmatrix}$$

Let $X = (x_1, \dots, x_n)$

Then $\text{Cov}(X)$

$$= \begin{pmatrix} \sigma_{11} & \sigma_{12} & \dots & \sigma_{1n} \\ \sigma_{21} & \sigma_{22} & \dots & \sigma_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{n1} & \sigma_{n2} & \dots & \sigma_{nn} \end{pmatrix}$$

$$\begin{pmatrix} \sigma_{21} & \sigma_{22} & \dots & \dots \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{n1} & \sigma_{n2} & \dots & \sigma_{nn} \end{pmatrix}$$

where

$$\sigma_{ii} = \text{Var}(X_i)$$

$$\sigma_{ij} = \text{Cov}(X_i, X_j)$$

Multivariate Normal
(Gaussian) RV

let $\underline{X}$ be a p-variate
RV with

$$E(\underline{X}) = E \begin{bmatrix} X_1 \\ \vdots \\ X_p \end{bmatrix} = \begin{bmatrix} E(X_1) \\ \vdots \\ E(X_p) \end{bmatrix} = \begin{pmatrix} \mu_1 \\ \vdots \\ \mu_p \end{pmatrix}$$

$$= \underline{\mu}$$

$$\text{Cov}(\underline{x}) = \underline{\Sigma}$$

if $\underline{x}$ has joint PDF

$$f_{\underline{x}}(\underline{x}) = \frac{1}{(2\pi)^{p/2} |\underline{\Sigma}|^{1/2}} e^{-\frac{1}{2}(\underline{x}-\underline{\mu})^T \underline{\Sigma}^{-1}(\underline{x}-\underline{\mu})}$$

for $\underline{x} \in \mathbb{R}^p$

then $\underline{x}$ is said to follow p -variate Gaussian distribution

with parameters μ & Σ . This is denoted by $\underline{x} \sim N_p(\underline{\mu}, \underline{\Sigma})$

If $p = 1$, the PDF reduces to

$$f_x(x) = \frac{1}{(2\pi)^{\frac{1}{2}}(\sigma^2)^{\frac{1}{2}}} e^{-\frac{1}{2}(x-\mu)(\sigma^2)^{-1}(x-\mu)}$$
$$= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

This is univariate $N(\mu, \sigma^2)$ PDF.

$\rightarrow$ σ^2 is a scalar matrix

Properties of Cov matrix

- (i) It is a symmetric matrix
- (ii) It is a positive semi definite matrix.
(without proof)

Hint: ① obvious

Lemma: If $\underline{X}$ is a
RV with $E(\underline{X}) = \underline{\mu}$

and $\text{Cov}(\underline{X}) = \underline{\Sigma}$, then
for $\underline{A} \in \mathbb{R}^{p \times p}$ & $\underline{b} \in \mathbb{R}^p$

$$(i) E(\underline{A}\underline{X} + \underline{b}) = \underline{A}\underline{\mu} + \underline{b}$$

$$(1) \text{ Cov}(\tilde{A}\tilde{x} + \tilde{b}) \\ = \tilde{A} \tilde{\Sigma} \tilde{A}^T$$

without proof

Lemma: For bivariate normal RV, zero covariance implies independence of components.

Hint:

$$\text{Cov}(\tilde{x}) = \begin{pmatrix} \sigma_1^2 & 0 \\ 0 & \sigma_2^2 \end{pmatrix}$$

$$f_{\tilde{x}}(x_1, x_2) = \frac{1}{2} \left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2$$

$$= \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{1}{2} \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2}$$

$$\frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{1}{2} \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2}$$

Identically distributed RVs

Two RVs X and Y are said to be identically distributed if they have same CDF.

(i) If X and Y are discrete then identically distributed requires same PMF

(ii) Continuous

PDF

Notation $X \stackrel{d}{=} Y$

Independently and
identically distributed
(IID)

Def: A set of RVs
 $X_1, \dots, X_n$ are said to
be IID if

$$(i) \quad X_i \stackrel{d}{=} X_j \quad \forall i \neq j$$

(ii) $(X_1, \dots, X_n)$ are
independence.

Theorem: Let $X_1, \dots, X_n$
are IID RVs with
mean $\mu = E(X_1)$ and
variance $\sigma^2 = \text{Var}(X_1)$.

Define
$$\bar{X}_n = \frac{X_1 + \dots + X_n}{n},$$

it is known as sample
mean.

Then,

$$(i) \quad E(\bar{X}_n) = \mu$$

$$(ii) \quad \text{Var}(\bar{X}_n) = \frac{\sigma^2}{n}$$

Moment Generating Function

(M4F)

Def: Let X be a RV.
The MGF of X is
denoted by $M_X(t)$, $t \in \mathbb{R}$,
and defined as

$$M_X(t) = E(\underline{e^{tx}})$$

$$E(e^{tx}) = 1 + t \underline{E(X)} + \frac{t^2}{2!} \underline{E(X^2)} + \dots$$
